The Size of Fiscal Multipliers and the Stance of Monetary Policy in Developing Economies

Jair N. Ojeda-Joya Oscar E. Guzman

Working Paper No. 106 June 2017

INSTITUTE OF EMPIRICAL ECONOMIC RESEARCH

Osnabrück University Rolandstraße 8 49069 Osnabrück Germany

The Size of Fiscal Multipliers and the Stance of Monetary Policy in Developing Economies

Jair N. Ojeda-Joya

Oscar E. Guzman*

Abstract

In this paper we estimate the effect of government consumption shocks on GDP using a panel of 21 developing economies. Our goal is to better understand the reasons for the low fiscal multipliers found in the literature by performing estimations for alternative exchange rate regimes, business-cycle phases, and monetary policy stances. In addition, we perform counterfactual simulations to analyze the possible gains from fiscal-monetary policy coordination. The results imply that government consumption shocks are usually followed by monetary policy tightening in developing economies with flexible regimes. Our simulations show that this reaction partially explains the presence of low fiscal multipliers in these economies. On the other hand, we find that government consumption shocks imply higher multipliers in developing economies during fixed regimes, economic booms or monetary expansions. In particular, implementing fiscal programs during monetary expansions seems to improve significantly their economic stimulus.

Keywords: Fiscal Policy, Monetary Policy, Structural Vector Autoregression, Exchange Rate Regime, Panel VAR

JEL classifications: E62, E63, F32

[•] This version: May 2017. The findings, recommendations, interpretations and conclusions expressed in this paper are those of the authors and do not necessarily reflect the view of the Central Bank of Colombia or of the Office of the National Comptroller. We are grateful to Frank Westermann, Jan Kuckuck, Hernando Vargas, Hernán Rincón and participants of the economics seminar at Osnabrück University for their generous comments and suggestions. We also thank the Swiss Government's Bilateral Assistance and Capacity Building for Central Banks (BCC) program for financial support and Gabriela Rubio for excellent research assistance.

[^] Corresponding author: Senior Research economist, Banco de la Republica. Cr. 7 No. 14-78, Bogotá D.C., Colombia. E-mail: jojedajo@banrep.gov.co

Economic Advisor, Office of the National Comptroller, Bogotá D.C., Colombia. E-mail: oscar.guzman@contraloria.gov.co

1. Introduction

Having reliable estimations of fiscal multipliers is exceptionally important for central governments when evaluating new fiscal programs. However, there is a notable result for developing economies: their fiscal multipliers are very small, (Ilzetzki et al, 2013). On the other hand, fiscal multipliers which are derived from SVAR models include the predicted future path of policy instruments (Kuckuck and Westermann, 2014). This is because such estimations are typically performed with data which captures the historical reaction of monetary and tax policies to government consumption shocks. In this paper, we further explore the size of fiscal multipliers in developing economies by empirically evaluating the role of the implied monetary policy reactions in their determination.

We study a panel of quarterly data and 21 developing economies¹. We use a Structural Vector Auto Regression (SVAR) (Blanchard and Perotti, 2002), within panel data estimation. In the case of fiscal policy, a key identification assumption is that changes in government consumption require at least a quarter to respond to innovations in the remaining macro variables, including GDP. In the case of monetary policy, we use the money stock as a control variable for money demand which reacts quickly to the macro environment. In contrast, policy interest rates take longer to react to macroeconomic news.

Opposite to the case of developed economies, our impulse-response functions imply that government consumption shocks are followed by monetary policy tightening in developing economies with flexible exchange rates. In addition, we perform counterfactual simulations to obtain fiscal multipliers which are free of the monetary policy reaction implied by the data. That is, we shut down the estimated policy rate reaction and allow all the remaining feedbacks of the system to operate. These simulations show that such tightening partially explains the presence of low multipliers in these developing economies. The results also imply that further fiscal-monetary policy coordination might be useful during the implementation of policy programs. In particular, we show that when the government consumption shock coincides with a monetary expansion, the multiplier increases significantly.

The related literature has found that the size of fiscal multipliers is a function of specific macroeconomic features of the economies under study. Favero et al (2011) find heterogeneous results for the size of fiscal multipliers within a sample of developed economies even after controlling for macroeconomic features and using the narrative method to identify fiscal shocks. Auerbach and Gorodnichenko (2012) show that positive multipliers are mostly associated to recessions according to their estimations using OECD data and local projection methods. Kraay (2012, 2014) finds positive multipliers (around 0.5) in developing economies using credits from multiplieral organizations as identification device. These multipliers are

_

¹ We also briefly study a panel of 20 developed economies to initially compare the size of fiscal multipliers and the monetary policy reaction. For comparison purposes, the whole set of 41 countries is similar to the one studied by Ilzetzki et al (2013).

found to be significantly lower in open economies but still positive. Corsetti et al (2012) find that multipliers are low in flexible exchange rate regimes and high during financial crises in developed economies.

There is also recent literature which focuses on developing economies. Ilzetzki et al (2013) find that fiscal multipliers are low in developing economies, especially, during flexible exchange rate regimes and when public debt is relatively high. In addition, Riera-Crichton et al (2015) discover that fiscal multipliers are larger during recessions, especially when they coincide with countercyclical fiscal policies. Hory (2016) studies a panel of emerging and advanced economies and finds that emerging economies can improve the size of fiscal multipliers by improving their government indebtedness indicators. Similarly, Huidrom et al (2016) use panel data for 34 economies (emerging and developed) to estimate fiscal multipliers for alternative fiscal positions. Their results show that multipliers are very small during episodes of high public debt and large fiscal deficits. In addition, the fiscal position is much more relevant than the business cycle to determine the size of these multipliers.

The literature on the interaction between monetary and fiscal policies is mostly based on general equilibrium models. Davig and Leeper (2011) use a DSGE model to study monetary-fiscal policy interaction and find that the results crucially depend on whether these polices have an active or passive role after macroeconomic shocks. On the other hand, Woodford (2011) shows, within a stylized macro model, that the size of fiscal multipliers crucially depends on the monetary policy reaction. These multipliers are even higher at the zero lower bound. Born et al (2013) also show, within an open new-Keynesian model, that the monetary regime is very important on the determination of the size of fiscal multipliers. Chinn (2013) reviews most of the theoretical approaches and concludes that empirical estimations of fiscal multipliers are crucial since models with alternative assumptions and calibrations lead to widely different GDP responses to fiscal shocks.

There are two recent contributions that empirically examine the effect of the monetary policy reaction on the size of fiscal multipliers. Belinga and Lonkeng (2015) find that the fiscal multiplier in the US is higher if it coincides with accommodative monetary policy. Similarly, Pyun and Rhee (2015) study a panel of 21 OECD economies during the aftermath of the Great Financial Crisis and find that the combination of fiscal and monetary expansions contributed to fiscal multipliers greater than 1. To the best of our knowledge, this work is the first to study the effect of monetary-fiscal policy interactions on the size of fiscal multipliers within a set of only developing economies.

The rest of this paper is organized as follows. Section 2 describes the data. Section 3 reports the econometric methodology. Section 4 shows the empirical results for alternative macroeconomic features of the countries under study. Finally, Section 5 makes some concluding statements.

2. Data Description

We use data for 21 developing economies: Argentina, Botswana, Brazil, Bulgaria, Chile, Colombia, Croatia, Ecuador, El Salvador, Hungary, Latvia, Lithuania, Malaysia, Mexico, Peru, Poland, Romania, Slovak Republic, South Africa, Thailand and Turkey. We choose this set of countries due to their availability of fiscal and monetary macro data². We follow the World Bank's classification to define the "developing" category.

For each country we retrieve data to compute the following quarterly indicators: i) real gross domestic product, ii) real government consumption, iii) real effective exchange rate, iv) real monetary policy rate, and v) real Money Balances (M2). Series for all countries span the first quarter of 2000 until the first quarter of 2015. These data have good quality and fairly easy access since they are retrieved from International Financial Statistics (IFS), OECD, Eurostat, national statistical agencies and central banks.

We compute annual growth rates for each variable in each country to correct for non-stationarity and seasonality of the data. That is, we compute the difference between the log level of each variable at t and t-4. In the case of the real interest rate we do not use logs but only differences between percentage points. In the Appendix, we present results of panel unit root tests for all five panels of variables, after differencing (Tables 2 and 3).

3. Econometric Methodology

The main challenge when computing fiscal multipliers is the proper identification of fiscal shocks. We use the Structural Vector Auto Regression (SVAR) approach first employed by Blanchard and Perotti (2002). In this approach, a key identification assumption is that policy decisions about government consumption require at least a quarter to respond to innovations in other macroeconomic variables, including GDP.

Since our focus is the interaction between monetary and fiscal policies, we need to simultaneously identify both shocks. To identify monetary policy shocks, we use the SVAR approach following Kim (2003) and Anzuini et al (2013). Their strategy consists of controlling for money demand and including short-run restrictions so that money demand contemporaneously react to interest rate and output shocks. Monetary policy rates instead, take longer to react to output innovations.

We use the Cholesky decomposition to perform this identification. The variables are included in the following ordering: Government consumption, monetary policy rate, GDP, real money

² Our dataset also includes the following 20 developed economies: Australia, Belgium, Canada, Denmark, Finland, France, Germany, Greece, Iceland, Ireland, Israel, Italy, Netherlands, Norway, Portugal, Slovenia, Spain, Sweden, United Kingdom, and United States of America.

balances (M2) and real exchange rate. Such ordering implies that the real exchange rate and M2 are the variables that react contemporaneously to shocks on the remaining variables of the system. In addition, policy variables are those with the most lagged reaction to macroeconomic shocks³.

Alternative identification methods for fiscal shocks use natural experiments such as military build-ups (Ramey and Shapiro, 1998) or large international loan disbursements (Kraay, 2012 and 2014). These data are not available for all the countries in our sample. In addition, strict assumptions about the unpredictability by the private sector of these government operations are required and might not hold in every economy.

We estimate the following Panel VAR system:

$$Y_{it} = Y_{it-1}A_1 + Y_{it-2}A_2 + \dots + Y_{it-p}A_pu_{it} + u_i + e_{it} \quad i = 1, \dots, N \quad t = 1, \dots, T$$
 (1)

This is a homogeneous panel VAR of order p, with k dependent variables and with panel-specific fixed effects. The vector Y_{it} contains the dependent variables for country i at quarter t. Country-specific fixed effects are shown in the k-vector u_i . Idiosyncratic errors for each country and quarter, are contained in the k-vector e_{it} . We want to estimate the $(k \times k)$ matrices of coefficients $A_1, A_2, \cdots A_p$. The errors are assumed to be zero-mean and serially uncorrelated.

We follow the estimation and lag selection procedures suggested by Abrigo and Love (2016). The optimal lag order (p) is selected using the consistent model selection criteria (MMSC) proposed by Andrews and Lu (2001) which are based on the test for over-identifying restrictions (J statistic) devised by Hansen (1982). The estimation is performed using a GMM approach to Equation (1). This method employs instruments to guarantee the consistency of estimators, especially in fixed T and large N settings. Instead of a first difference transformation, this method uses forward orthogonal transformation which minimizes data loss when instrumenting. Finally, the whole system is estimated as a system of equations since it allows efficiency gains and makes cross-equation testing possible.

We perform counterfactual simulations of the effects of government consumption shocks by shutting down the estimated monetary policy reaction and allowing all the remaining feedbacks of the system to operate. These exercises allow obtaining impulse-response functions which are free of the monetary policy reaction implied by the data. Therefore, the difference between unrestricted and counterfactual fiscal multipliers can be used to compute the approximate gains (or losses in some cases) from a coordinated monetary-fiscal policy in terms of GDP.

We compute fiscal multipliers from 0 to 20 quarters using the following standard formula:

$$fm(T) = \frac{\sum_{t=0}^{T} (1+r)^{-t} \Delta y_t}{\sum_{t=0}^{T} (1+r)^{-t} \Delta g_t}$$
 (2)

³ Mountford and Uhlig (2009) propose an alternative identification approach for fiscal policy, monetary policy and business-cycle shocks within the same VAR system. This approach uses sign and orthogonality restrictions for the identification of shocks. However, their results for the US are in line with other estimations.

In this equation, the cumulative fiscal multiplier in period T is defined as the present value of the implied changes in GDP divided by the present value of changes in government consumption. Both Δy_t and Δg_t are extracted from their respective impulse-response functions. The discount interest rate (r) corresponds to the average real policy rate across all countries and periods of analysis. Impact, medium-run and long-run multipliers correspond to T=0, T=10 and T=20, respectively, in Equation 2. This definition of fiscal multipliers is standard in the literature, except for the use of a real interest rate to discount future flows of GDP and government consumption (Chinn, 2013). However, for low values of this interest rate, the effect of discounting on the size of the multiplier is very small.

4. Empirical Results

In this section, we initially present results on the implications of a government consumption shock in developed and developing economies. Then, we describe the results during alternative exchange rate regimes, business-cycle phases and monetary policy-stances in developing economies. In every case, we show impulse-response functions for all variables of the system and then the implied cumulative fiscal multiplier. Every figure compares predicted and counterfactual multipliers. Using the procedure by Andrews and Lu (2001), optimal lag orders for developed and developing economies are p = 1 and p = 2, respectively. See tables 4 and 5 in the Appendix.

4.1 Developed versus Developing Economies

Figure 1 shows impulse responses of all variables in the system to a one standard-deviation government consumption shock in high-income economies. These graphs describe both the unrestricted and counterfactual responses during 20 quarters after the initial shock. The counterfactual responses exclude the monetary policy reaction implied by the data.

It is clear from Figure 1 that the increase of government consumption has a consistent positive effect on GDP lasting more than 20 quarters. Notice most of the monetary policy reaction is expansionary since interest rates are reduced during all 20 quarters, except for the impact effect. However, the monetary policy effect on the transmission of the fiscal shock on GDP is very small since the counterfactual response is very similar to the unrestricted one.

Figure 2 presents the cumulative fiscal multiplier implied by the GDP response depicted in Figure 1. The impact multiplier (immediately after the shock) is 0.83. It starts gradually increasing until reaching 2.8 in the medium-run (10th quarter) and 3.05 in the long run (20th quarter). The counterfactual multiplier is slightly lower (3.04 in the long-run). Therefore, there is no role for further fiscal-monetary policy coordination in developed economies.

Figure 1. Impulse-response functions in developed economies. one standard-deviation government consumption shock

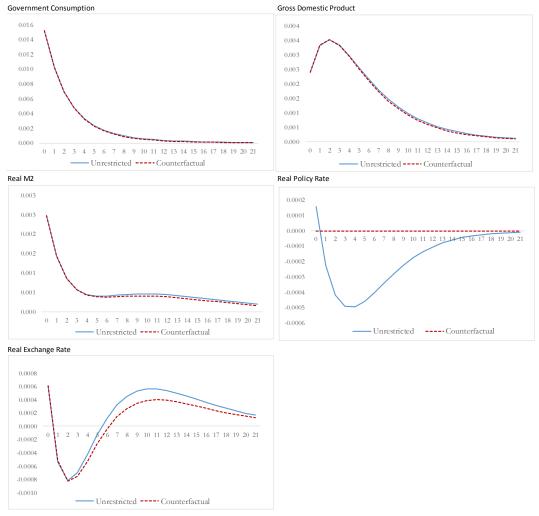


Figure 2. Cumulative fiscal multiplier in developed economies.

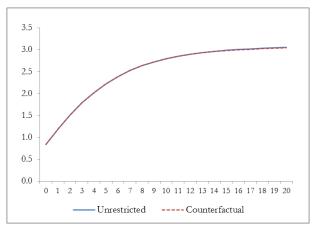


Figure 3. Impulse-response functions in developing economies. one standard-deviation government consumption shock

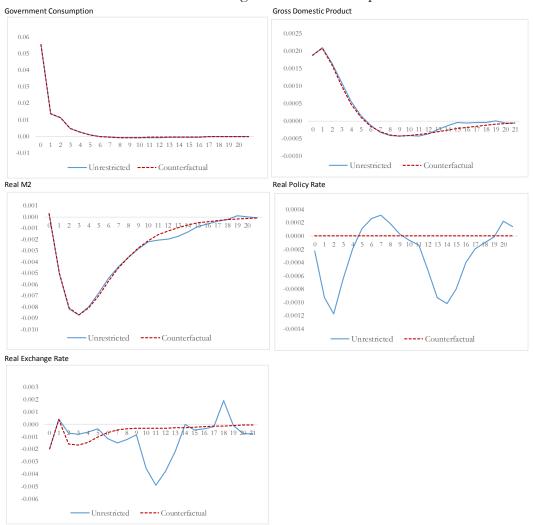


Figure 4. Cumulative fiscal multiplier in developing economies.

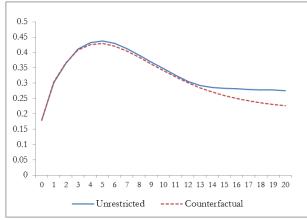


Figure 3 shows impulse responses for developing economies. The response of GDP is initially positive during the first 6 quarters after the shock. Then it becomes negative but close to zero during the remaining quarters. Notice that the monetary policy reaction is cyclical, initially expansionary and then, 4 to 9 quarters after the shock, it becomes contractive. At the end, it becomes again expansionary. The counterfactual GDP response is slightly below the unrestricted response.

The impact multiplier is 0.18 and its maximum is 0.44 which occurs 5 quarters after the shock, see Figure 4. Afterwards, it starts decreasing gradually until reaching 0.35 in the medium run and 0.28 in the long-run. The counterfactual multiplier is slightly lower which shows that further fiscal-monetary policy coordination is no needed for the average developing economy. In the following figures we further explore these results in developing economies by exploring alternative episodes of exchange-rate regimes, business-cycle phases and monetary policy stances.

4.2 Exchange-Rate Regimes in Developing Economies:

We study whether low fiscal multipliers in developing economies can be explained by the exchange rate regime. Following Ilzetzki et al (2013), the dataset is divided into episodes of flexible and fixed regimes using the classification by Reinhart and Rogoff (2004) updated with information from the International Monetary Fund (IMF) until 2015. Fixed-rates episodes are defined by legal tenders, hard pegs, crawling pegs, and bands. All other cases are classified as flexible regimes. Most countries alternate between fixed and flexible exchange rate regimes.

In the case of fixed regimes, the response of GDP is positive during 6 quarters after the shock and then becomes essentially zero. In addition, monetary policy becomes consistently expansionary during all 20 quarters after the shock; see Figure 11 in the Appendix. The fiscal multiplier (Figure 5) is initially 0.34 on impact, and reaches a maximum of 0.53, 5 quarters later. The medium-run and long-run multipliers are 0.47 and 0.5, respectively. The counterfactual simulation implies that there is no role for further fiscal-monetary policy coordination. These results are consistent with the predictions of the Mundell-Fleming model for fixed exchange regimes.

Figure 5. Cumulative fiscal multiplier in developing economies with fixed exchange rate.

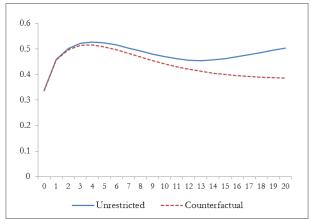
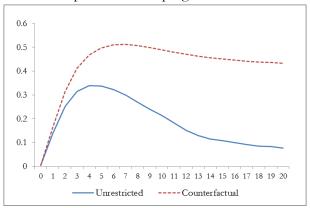


Figure 6. Cumulative fiscal multiplier in developing economies with flexible exchange rates.



Source: Authors' computations

In the case of flexible regimes, the unrestricted GDP response is positive during 6 quarters and then it becomes negative but close to zero (Figure 12). This reaction is partly related to the increase in the monetary policy rate observed in the same figure since the counterfactual GDP response is clearly higher. Figure 6 shows that the implied fiscal multiplier is essentially zero on impact and then it gradually grows to reach a maximum of 0.34, 4 quarters after the shock. Medium and long-run unrestricted multipliers are 0.21 and 0.08, respectively. The counterfactual simulation shows that in this case there is a key role for fiscal-monetary policy coordination since counterfactual multipliers are much higher: 0.49 and 0.43, in the medium and long-run, respectively. This result is in line with the predictions of standard open-economy models in which the monetary policy reaction is the reason for low fiscal multipliers in economies with flexible regimes; see for example, McCallum (1996). However, the size gap between developed and developing economies is not fully explained by this lack of coordination.

4.3 Business-Cycle Phases in Developing Economies

In this section we divide the set of developing economies into two alternative stances of economic activity: booms and recessions. We would like to explore whether the low fiscal multipliers in these economies are related to the business-cycle phase. This exercise follows the approach by Auerbach and Gorodnichenko (2012), Riera-Crichton et al (2015), among other works.

We apply a quarterly Hodrick-Prescott filter to real GDP data, country by country, to estimate their trend. We define as recessions (booms) those periods when observed GDP is below (above) such trend⁴. Notice that all countries have several alternative episodes of booms and recessions and our econometric methodology allows computing fiscal multipliers on these subsets of data. Since the average duration of these phases is around 8 quarters in developing economies, we are presenting impulse responses only up to 10 quarters after the shock.

In the case of fiscal shocks during GDP booms, the GDP response is positive during most periods after the shock. Similarly, we observe an increase of the monetary policy rate during all periods, except the first quarter (Figure 13). The implied cumulative fiscal multiplier (Figure 7) is positive on impact (0.22) and reaches a maximum of 0.51 in the sixth quarter. Then it starts decreasing gradually until reaching a medium-run level of 0.47. There is only a very small role for monetary-fiscal policy coordination since the counterfactual multiplier is only slightly higher than the unrestricted one.

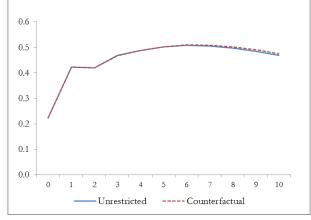


Figure 7. Cumulative fiscal multiplier in developing economies during economic booms.

Source: Authors' computations

10

⁴ In practice, since the optimal lag order is 2, we only consider boom or recession episodes of at least 3 quarters.

0.30 0.25 -0.20 -0.15 -0.10 -

Unrestricted

Figure 8. Cumulative fiscal multiplier in developing economies during economic recessions.

Source: Authors' computations

0.05 0.00 -0.05

In the case of economic recessions, the unrestricted GDP response is initially zero and then remains positive from quarters 1 to 5. In contrast, the monetary policy response is cyclical, remaining positive only between quarters 2 and 6 (Figure 14). Therefore, the implied fiscal multiplier (Figure 8) is initially close to zero but then it starts increasing until reaching 0.26 in the fifth quarter and a very similar level in the medium-run (tenth quarter). There is a small role for fiscal-monetary policy coordination since the multiplier is slightly higher during the countercyclical simulation.

---- Counterfactual

We find that cumulative multipliers are higher during booms than during recessions. This result contrasts with recent literature, for example, Auerbach and Gorodnichenko (2012). It is important to point out that while most of the literature works with datasets from developed economies or mostly developed economies (OECD data), we study a dataset of only developing economies. Further works on the reasons for this reversal in the relation between business cycle and fiscal multiplier are needed. A first hypothesis would be that during recessions, developed economies have much more fiscal space than developing ones.

4.4 Monetary Policy Stances in Developing Economies

An additional exercise is studying whether the transmission of fiscal policy is affected by the current stance of monetary policy. Woodford (2011), for example, shows a model in which fiscal policy is more effective when monetary policy approaches the zero lower bound. For this exercise, first we compute Hodrick-Prescott trends on the real policy rate, country by country. This trend would be a proxy of the natural interest rate. We define episodes of monetary contraction (expansion) as those quarters when the real policy rate is above (below) its trend⁵.

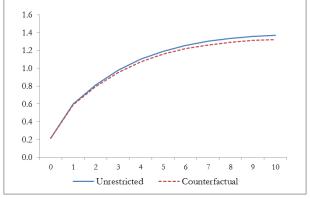
⁵ Similarly to business-cycle phases, in practice we only consider monetary-policy stances of at least 3 quarters long since the Panel VAR includes 2 lags for all variables.

Since these monetary policy stances have similar durations to business cycles, we show impulse responses and cumulative multipliers only up to 10 quarters after the shock.

In the case of monetary expansions, the GDP response is positive during all 10 quarters after the shock. In addition, the policy rate response is negative and the real exchange rate tends to depreciate. Notice that both reactions are expected to improve the transmission of the fiscal policy shock, (Figure 15). The implied cumulative multiplier (Figure 9) is small on impact (0.22); then, it starts to increase gradually and reaches a medium-run level of 1.37 which is much higher than the average multiplier estimated for developing economies (Figure 4). It seems that, similarly to Woodford (2011), the relatively low real interest rates bring about good financial conditions that increase the effect of the fiscal stimulus.

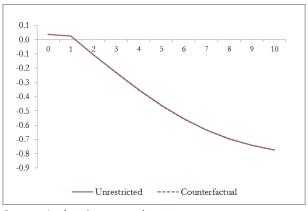
Figure 9. Cumulative fiscal multiplier in developing economies during monetary expansions.

1.6



Source: Authors' computations

Figure 10. Cumulative fiscal multiplier in developing economies during monetary contractions.



During monetary contractions, the GDP reaction is negative starting in the second quarter after the shock. Meanwhile, the implied monetary policy reaction is cyclical and the effect on the real exchange rate is very small (Figure 16). This GDP reaction leads to an insignificant impact multiplier (Figure 10) of 0.04 that starts going downward until reaching -0.78 in the medium run. The counterfactual multiplier is very similar. This negative multiplier is likely the result of increased sovereign risk in these economies due to the tighter financial conditions implied by the contractive monetary policy. However, these negative multipliers are not statistically different to zero as we show in Section 4.5.

This result is new and has important policy implications for developing economies; namely, fiscal stimulus programs are especially effective if implemented during monetary expansions. The mechanism is related to those explained by Woodford (2011) or Davig and Leeper (2011) on the higher effectiveness of fiscal expansions during loose financial conditions. This finding also shows the possibility of further fiscal-monetary policy coordination to improve the effects of fiscal stimulus programs.

4.5 Statistical Significance of Fiscal Multipliers

We perform statistical significance tests on fiscal multipliers by computing confidence bands on the cumulative GDP response. This estimation is based on Abrigo and Love (2016) and uses 10000 Montecarlo draws from the estimated panel VAR using Gaussian approximation. In Table 1, we present selected multipliers for alternative panel subsets and report whether they are significant with a confidence degree of at least 90%.

The results discussed previously in section 4 remain similar once the statistical significance results are taken into account. First, fiscal multipliers for developed economies are high and significant for all horizons. In contrast, multipliers for developing economies are low and only significant on impact for fixed exchange rate regimes, boom episodes and monetary expansions. On the medium-term, multipliers are only significant in developing economies if the shock takes place during a monetary expansion. The latter result is new and has interesting policy implications.

Table 1 – Summary of Unrestricted Fiscal Multipliers for Alternative Horizons

Panel Subset	Impact	5th quarter	Medium Term	Long Term	# of observations
	1	1	(10th quarter)	(20th quarter)	
Developed	0.83**	2.22**	2.80**	3.05**	1180
Developing	0.18**	0.44	0.35	0.28	1218
Developing – Fixed Reg.	0.34**	0.52	0.47	0.50	643
Developing – Flex. Reg.	0.00	0.34	0.21	0.08	605
Developing – Boom	0.22**	0.50	0.47	NA	328
Developing – Recession	-0.01	0.26	0.26	NA	445
Developing – Mon. Exp.	0.22**	1.19**	1.37**	NA	369
Developing – Mon. Con.	0.04	-0.46	-0.78	NA	371

^{**} Significant with a confidence degree of at least 90%; NA: Not available

5. Concluding Statements

In this work, we study the size of fiscal multipliers and the transmission channels of government consumption shocks, using panel data for 21 developing economies. The goal is to better understand the reasons for the existence of low fiscal multipliers in these economies. We also explore the role of fiscal-monetary policy coordination by computing counterfactual simulations in which the implied monetary policy reaction is elapsed. Furthermore, we estimate fiscal multipliers in subsets of the data for alternative degrees of development, exchange-rate regimes, business-cycle phases and monetary policy stances.

We find that relative to developed economies, fiscal multipliers in developing economies are positive but much lower. Among developing ones, we find higher multipliers in fixed exchange rate regimes, during economic booms and during monetary expansions. Fiscal multipliers are especially big during episodes of monetary expansion since they reach levels above one in the medium-run.

Using counterfactual simulations, we find that by enhancing fiscal-monetary policy coordination in developing economies with flexible regime, it is possible to increase the effects of fiscal policy. Namely, fiscal policy programs would me more effective if the monetary policy reaction was more gradual.

Our empirical results have important implications for developing economies. First, policy makers in these economies should be very careful before engaging in countercyclical fiscal expansions. These fiscal programs seem to have higher effects during economic booms and more importantly, during monetary expansions. This effectiveness can be further improved if monetary authorities delay interest rate increases during their implementation.

The transmission channels of fiscal policy should be further explored using empirical estimations as well as general equilibrium macro models. Our estimations suggest that the financial stability channel might be important since government consumption shocks are mostly followed by real money demand reductions in developing economies, in contrast to developed economies. Other channels not considered in this work are: sovereign default risk and political risk.

References

Abrigo, M. R., and I. Love, (2016). "Estimation of Panel Vector Autoregression in Stata," *The Stata Journal*, 16(3), pp. 778-804.

Andrews, D. W. K., and B. Lu (2001). "Consistent Model and Moment Selection Procedures for GMM Estimation with Application to Dynamic Panel Data Models," *Journal of Econometrics*, 101(1), pp. 123-164.

Anzuini, A., M. Lombardi, and P. Pagano, (2013). "The Impact of Monetary Policy Shocks on Commodity Prices," *International Journal of Central Banking*, 9(3), pp.119-144.

Auerbach, A., and Y. Gorodnichenko, (2012). "Fiscal Multipliers in Recession and Expansion," in: "Fiscal Policy after the Financial Crisis", National Bureau of Economic Research, Cambridge, MA, pages 63-98.

Belinga, V. and Lonkeng, C. (2015) "(Not) Dancing Together: Monetary Policy Stance and the Government Spending Multiplier". IMF Working Paper 115/14.

Blanchard, O., and R. Perotti, (2002). "An Empirical Characterization of the Dynamic Effects of Changes in Government Spending and Taxes on Output," *Quarterly Journal of Economics*, 117, pp. 1329-1368.

Born, B., F. Juessen, and G. Müller, (2013). "Exchange Rate Regimes and Fiscal Multipliers," *Journal of Economic Dynamics and Control*, 37, pp. 446-465

Chinn, M. (2013). "Fiscal Multipliers," The New Palgrave Dictionary of Economics, online edition.

Corsetti, G., A. Meier, and G. Müller, (2012) "What Determines Government Spending Multipliers?" *Economic Policy*, 72, October.

Davig, T., and E. M. Leeper, (2011). "Monetary-Fiscal Policy Interactions and Fiscal Stimulus," *European Economic Review*, 55(2), pp. 211-227.

Favero, C., F. Giavazzi, and J. Perego, (2011). "Country Heterogeneity and the International Evidence on the Effects of Fiscal Policy," *IMF Economic Review*, 59(4), pp. 652-682.

Hansen, L. P., (1982)."Large Sample Properties of Generalized Method of Moments Estimator," *Econometrica*, 50(4), pp. 1029-1054.

Hory, M. P., (2016). "Fiscal Multipliers in Emerging Market Economies: Can We Learn Something from Advanced Economies?" *International Economics*, 146, pp. 59-84

Huidrom, R., Kose, M., Lim, J., and Ohnsorge, F. (2016) "Do Fiscal Multipliers Depend on Fiscal Positions?" Policy Research Working Paper No. 7724, World Bank. June

Ilzetzki, E., E. Mendoza, and C. Végh, (2013). "How big (small?) are fiscal multipliers?," *Journal of Monetary Economics*, 60(2), pp. 239-254.

Im, K., M. Pesaran, and Y. Shin, (2003). "Testing for Unit Roots in Heterogeneous Panels," *Journal of Econometrics*, 115(1): pp. 53-74

Kim, S., (2003). "Monetary policy, foreign exchange intervention, and the exchange rate in a unifying framework," *Journal of International Economics*, 60, pp. 355-386.

Kraay, A., (2014). "Government Spending Multipliers in Developing Countries: Evidence from Lending by Official Creditors," *American Economic Journal: Macroeconomics*, 6(4), pp. 170-208.

Kraay, A., (2012). "How large is the Government Spending Multiplier? Evidence from World Bank Lending," *The Quarterly Journal of Economics*, 127(2), pp. 829-887.

Kuckuck, J., and F. Westermann, (2014). "On the size of fiscal multipliers: A counterfactual analysis," *Economics Letters*, 123(1), pp. 26-32.

McCallum, B. (1996). "International Monetary Economics". Oxford University Press, 280 p.

Pyun, J. and Rhee, D. (2015). "Fiscal Multipliers during the global financial crisis: fiscal and monetary interaction matters". *Contemporary Economic Policy*: 33 (1), pp. 207-220.

Ramey, V., and M. Shapiro, (1998). "Costly Capital Reallocation and the Effects of Government Spending," In: *Carnegie-Rochester Conference Series on Public Policy*, 46, pp. 145-194.

Reinhart, C. and K. Rogoff, (2004). "The Modern History of Exchange Rate Arrangements: A Reinterpretation," *The Quarterly Journal of Economics*, 119(1), pp. 1-48.

Riera-Crichton, D., C. Vegh, and G. Vuletin, (2015). "Procyclical and countercyclical fiscal multipliers: Evidence from OECD countries," *Journal of International Money and Finance*, 52, pp. 15-31.

Woodford, M., (2011). "Simple Analytics of Government Expenditure Multiplier," *American Economic Journal: Macroeconomics*, 3(1): pp. 1-35.

Appendix

Table 2: Panel Unit Root Tests – Developed

Test by Im, Pesaran and Shin (2003)

Null: Unit root (assumes common unit root process)

Panel	t-bar	P-value
Gov. Consumption	-2.8217	0.0000
GDP	-2.7561	0.0000
Policy rate	-3.0407	0.0000
Real exchange rate	-2.5220	0.0000
M2	-2.2732	0.0001

Source: Authors' calculations

Table 3: Panel Unit Root Tests – Developing

Test by Im, Pesaran and Shin (2003)

Null: Unit root (assumes common unit root process)

Panel	t-bar	P-value
Gov. Consumption	-4.0455	0.0000
GDP	-2.4721	0.0000
Policy rate	-3.6118	0.0000
Real exchange rate	-2.3654	0.0000
M2	-2.9201	0.0000

Source: Authors' calculations

Table 4: Lag Selection Criteria – Developed

\mathcal{O}			
Lags	MBIC	MAIC	MQIC
1	-476.36	77.70	-132.48
2	-419.96	54.95	-125.21
3	-367.38	28.38	-121.75
4	-283.14	33.47	-86.63

Source: Authors' calculations

Table 5: Lag Selection Criteria – Developing

Lags	MBIC	MAIC	MQIC
1	-549.04	-10.49	-201.26
2	-510.70	-31.11	-212.60
3	-420.22	-20.56	-171.81
4	-342.15	-22.42	-143.42

Source: Authors' calculations

Figure 11. Impulse-response functions in developing economies during fixed exchange rates

One standard-deviation government consumption shock

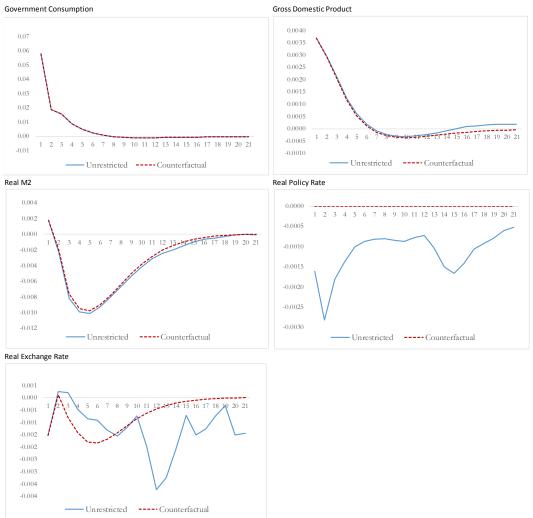


Figure 12. Impulse-response functions in developing economies during flexible exchange rates

One standard-deviation government consumption shock

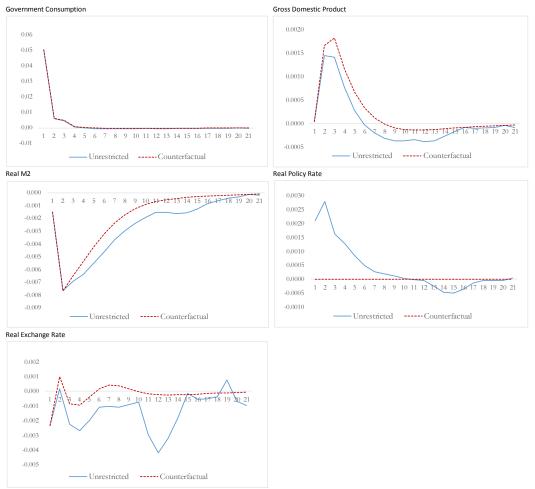


Figure 13. Impulse-response functions in developing economies during economic booms

One standard-deviation government consumption shock

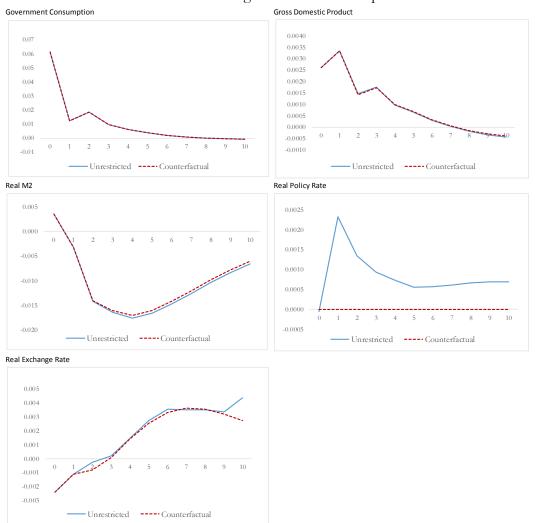


Figure 14. Impulse-response functions in developing economies during economic recessions

One standard-deviation government consumption shock

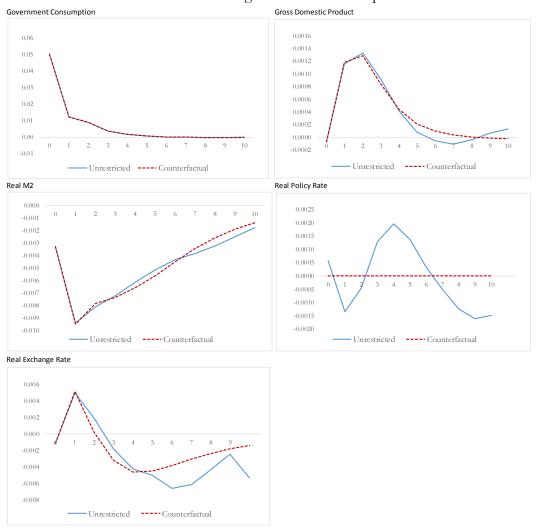


Figure 15. Impulse-response functions in developing economies during monetary expansions

One standard-deviation government consumption shock

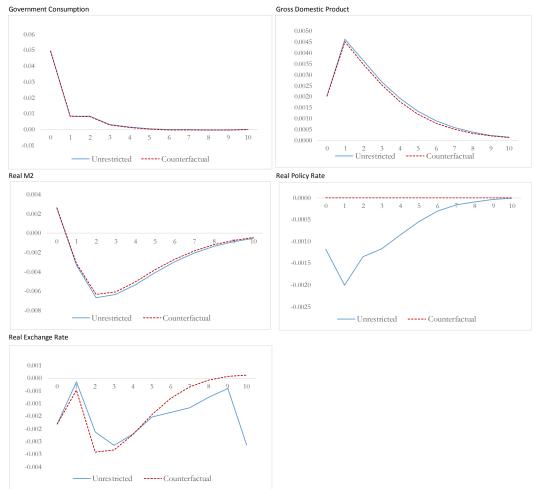


Figure 16. Impulse-response functions in developing economies during monetary contractions

One standard-deviation government consumption shock

